

# Cluster Expansion Method for Evolving Weighted Networks Having Vector-Like Nodes

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The cluster variation method known in statistical mechanics and condensed matter is revived for weighted bipartite networks. The decomposition (or expansion) of a Hamiltonian through a finite number of components, whence serving to define variable clusters, is recalled. As an illustration the network built from data representing correlations between (4) macro-economic features, i.e. the so-called *vector components*, of 15 EU countries, as (function) nodes, is discussed. We show that statistical physics principles, like the maximum entropy criterion points to clusters, here in a (4) variable phase space: Gross Domestic Product, Final Consumption Expenditure, Gross Capital Formation and Net Exports. It is observed that the *maximum* entropy corresponds to a cluster which does *not* explicitly include the Gross Domestic Product but only the other (3) “axes”, i.e. consumption, investment and trade components. On the other hand, the *minimal* entropy clustering scheme is obtained from a coupling necessarily including Gross Domestic Product and Final Consumption Expenditure. The results confirm intuitive economic theory and practice expectations at least as regards geographical connexions. The technique can of course be applied to many other cases in the physics of socio-economy networks.

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## 1. Introduction

In physics one is often interested about models with a finite number  $N$  of degrees of freedom, hereby denoted by  $\mathbf{s} = (s_1, s_2, \dots, s_N)$ , taking sometimes *discrete* values, in contrast to continuous ones, as in field theories. For instance, the variables  $s_i$  could take values [0 or 1] (binary variables), [-1,+1] (Ising spins), or [1, 2, ...  $q$ ] (Potts variables). Network nodes and/or links can possess such degrees of freedom which indicate the role of a few variables for characterizing or tying

nodes together; these variables serve, e.g., to be exemplifying clusters, communities, ... in the network. Several network characterization techniques based on related discrete value algebra exist in the literature [1, 2].

Let us recall that statistical mechanical models are defined through an energy function, like a Hamiltonian,  $\mathcal{H} = \mathcal{H}(s)$ ; the corresponding probability distribution at thermal equilibrium is the Boltzmann distribution

$$p(s) = \frac{1}{\mathcal{Z}} \exp(-\mathcal{H}(s)), \quad (1)$$

where the inverse temperature  $\beta = (k_B T)^{-1}$  has been absorbed into the Hamiltonian as often conventionally done;

$$\mathcal{Z} = \exp(-\mathcal{F}) = \sum_s \exp(-\mathcal{H}(s)) \quad (2)$$

is called the partition function and  $\mathcal{F}$  the free energy. The Hamiltonian is typically a sum of terms, each involving a small number of variables.

A technique which has been of interest a long time ago in condensed matter is the cluster variation approximation method [3–5]. The free energy or the Hamiltonian is expanded through a series in the variables by a systematic projection in order to define the interaction energy at each successive cluster size level. We re-introduce the technique here, suggesting its power for discussing network properties. The technique appears to be very general and could be useful to sort out features not observed otherwise. Whence for better framing the concept and rejuvenating the vocabulary we re-outline some theoretical consideration and rewrite well known formulae.

We take as an example and for illustration a finite size network, one made of nodes being EU countries characterized by their most usual (macroeconomic) features. The fluctuation correlations between these serve to define the so-called adjacency matrix, whence the *weights* of the links of the network. Some “discussion section” will indicate the interest of projecting the Hamiltonian, in an appropriate phase space in order to observe features. One property being measured will be the system entropy which is clearly a mapping of  $p(s)$  into macroscopic-like features based on economic indicators.

## 2. Theoretical considerations

A useful representation of interacting units (spins, agents, fields, ...) is given by the *factor graph*. A factor graph [6] is a bipartite graph made of variable nodes  $i, j, \dots$ , one for each variable, and function nodes  $a, b, \dots$ , one for each interaction term of the Hamiltonian. A link joins a variable node  $i$  and a function node  $a$  if and only if  $i \in a$ , that is the variable  $s_i$  appears in  $H_a$ , the term of the Hamiltonian associated to  $a$ . The Hamiltonian can then be written as

$$\mathcal{H} = \sum_a^N \mathcal{H}_a(s_a) \quad (3)$$

with  $s_a \equiv \{s_i, i \in a\} \dots$ . This sort of writing through the decomposition of a

Hamiltonian into terms describing clusters of different (increasing) sizes and functionalities has been shown to be of great interest, see [7], e.g. when one applies renormalization group-like techniques.

In combinatorial optimization problems, the Hamiltonian plays the role of a *cost function* and one is often interested in the low temperature limit  $T \rightarrow 0$ , where only minimal energy states (ground states) have a nonvanishing probability.

Probabilistic graphical models are usually defined in a slightly different way [8]. For example, in the case of *Markov random fields*, also called *Markov networks*, the joint distribution over all variables is given by

$$p(s) = \frac{1}{\widehat{\mathcal{Z}}} \prod_a \psi_a(s_a), \quad (4)$$

where  $\psi_a$  is called the potential, and

$$\widehat{\mathcal{Z}} = \sum_s \prod_a \psi_a(s_a). \quad (5)$$

Of course, a statistical mechanical model described by the Hamiltonian (3), corresponds to a probabilistic graphical model with potentials  $\psi_a = \exp(-\mathcal{H}_a)$ , and corresponding  $\mathcal{Z} = \widehat{\mathcal{Z}}$  and  $\mathcal{F} = \widehat{\mathcal{F}}$ .

Next we define a cluster  $\alpha$  as a subset of the factor graph such that if a function node belongs to  $\alpha$ , then all the variable nodes  $s_\alpha$  also belong to  $\alpha$ ; we notice that the converse needs not be true, otherwise the only legitimate clusters would be the connected components of the factor graph. Given a cluster we can write its probability distribution, defined as the ratio between the number of realized connections and the number of all possible connections, as

$$p_\alpha(s_\alpha) = \sum_{s \in \alpha} p(s) \quad (6)$$

and its entropy

$$\mathcal{S}_\alpha(s_\alpha) = - \sum_{s \in \alpha} p(s) \ln p(s). \quad (7)$$

### 3. Illustration

As a short illustration, let us consider the function nodes to be countries and the variables to be macroeconomic indicators [9], i.e.:

1. Consider the nodes to be the first (in time) 15 EU countries. Let the country names be abbreviated according to the Roots Web Surname List (RSL) [10] which uses 3 letters standardized abbreviations.
2. Suppose that we are interested in a vector describing each country (Hamiltonian or) “thermodynamic state” with 4 components, i.e.  $s_1 \equiv$  Gross Domestic Product (GDP),  $s_2 \equiv$  Final Consumption Expenditure (FCE),  $s_3 \equiv$  Gross Capital Formation (GCF), and  $s_4 \equiv$  Net Exports (NEX). The World Bank database [11] is here used as data source. Let the data be taken from 1994 to 2004 for GDP and from 1994 to 2003 for FCE, GCF and NEX, respectively.

The yearly fluctuations of these four variables are easily calculated and their auto- and cross-correlation matrices easily obtained; see e.g. a discussion for GDP in [12, 13] and more detail elsewhere [14]. Essentially, the correlations can be calculated for a time window of given size moving along the time axis; these are used for getting the statistical distances among countries, e.g.  $A$  and  $B$ , for various time window sizes  $T$  at various times  $t$ , where  $t$  is the final point of the interval, i.e.

$$d_s(A, B)_{(t,T)} = \sqrt{2[1 - C_{(t,T)}(A, B)]} \quad (8)$$

as in [15] where

$$C_{(t,T)}(A, B) = \frac{\langle AB \rangle_{(t,T)} - \langle A \rangle_{(t,T)} \langle B \rangle_{(t,T)}}{\sqrt{(\langle A^2 \rangle_{(t,T)} - \langle A \rangle_{(t,T)}^2)(\langle B^2 \rangle_{(t,T)} - \langle B \rangle_{(t,T)}^2)}}. \quad (9)$$

The brackets  $\langle \dots \rangle$  denote the expectation value of the “ $A, B$  time series”, here in the interval  $(t - T, t)$ .

These distances are thus mapped onto ultrametrical distances, as in the classical minimum spanning tree (MST) method. By calculating the statistical distances *with respect to the average value of the index* (seen here as for an “average” country), we get a country hierarchy that proves to be changing from a time interval to another when the (constant size) time window is moved over the full time span. The correlation coefficients refer to the movement of the countries inside this hierarchy.

In order to exemplify this method, the corresponding steps for  $s_2 \equiv$  FCE are explicitly shown below (for  $s_1 \equiv$  GDP the first steps are explicitly described in [16]). After the (virtual) “average” country is introduced in the system, the statistical distances corresponding to the fixed 5 years moving time window can be calculated and set in increasing order. The minimal path length (MPL) connections to the “average” country can be established for each country in every time interval (Table I). The resulting hierarchy is readily found to be changing from a time interval to another. The above procedure is repeated for each macroeconomic indicator, leading to similar three Tables to Table I. Next, a time independent “correlation matrix” can be built, at this stage for the *country movement fluctuations inside the hierarchy*, i.e. averaging the relative

TABLE I  
The MPL distances to the “average country”. Indicator: FCE ( $\equiv s_2$ ). The moving time window size is  $T = 5$  years for the data [11] taken from 1994 to 2003.

	aut	bel	deu	dnk	esp	fin	fra	gbr	grc	irl	ita	lux	nld	prt	swe
94–98	0.88	0.65	0.85	0.88	0.65	0.37	0.65	0.65	0.65	0.65	0.37	0.65	0.65	0.65	0.65
95–99	0.79	0.79	0.79	0.81	0.79	0.41	0.79	0.79	0.93	0.79	0.53	0.59	0.79	0.79	0.79
96–00	1.02	1.02	1.02	1.02	1.02	1.02	1.02	1.02	1.02	1.02	0.26	1.02	1.02	1.02	1.02
97–01	0.51	0.51	0.51	0.65	0.51	0.73	0.88	0.51	0.65	0.51	0.33	0.88	0.51	0.51	0.51
98–02	0.52	0.52	0.52	0.96	0.52	0.66	0.95	0.65	0.96	0.52	0.35	1.19	0.52	0.52	0.52
99–03	0.45	0.42	0.45	1.00	0.45	0.53	0.40	0.46	1.00	0.42	0.30	0.92	0.45	0.45	0.45

TABLE II

The correlation matrix of EU-15 country movements inside the hierarchy. Indicator: FCE. The moving time window size is 5 years for data taken from 1994 to 2003.

	aut	bel	deu	dnk	esp	fin	fra	gbr	grc	irl	ita	lux	nld	prt	swe
aut	1.00	<b>0.92</b>	<b>1.00</b>	0.23	<b>0.92</b>	0.21	0.38	0.87	0.03	<b>0.92</b>	0.07	-.34	<b>0.92</b>	<b>0.92</b>	<b>0.92</b>
bel		1.00	<b>0.94</b>	0.23	<b>1.00</b>	0.45	0.56	<b>0.97</b>	0.28	<b>1.00</b>	0.06	-.15	<b>1.00</b>	<b>1.00</b>	<b>1.00</b>
deu			1.00	0.24	<b>0.93</b>	0.24	0.40	0.89	0.07	<b>0.94</b>	0.07	-.32	<b>0.93</b>	<b>0.93</b>	<b>0.93</b>
dnk				1.00	0.26	0.22	0.14	0.35	0.75	<b>0.23</b>	-.41	0.44	0.26	0.26	0.26
esp					1.00	0.45	0.53	0.97	0.31	<b>1.00</b>	0.04	-.15	<b>1.00</b>	<b>1.00</b>	<b>1.00</b>
fin						1.00	0.65	0.49	0.34	0.45	0.68	0.68	0.45	0.45	0.45
fra							1.00	0.64	0.05	0.56	-.05	0.38	0.53	0.53	0.53
gbr								1.00	0.40	<b>0.97</b>	0.03	0.02	<b>0.97</b>	<b>0.97</b>	<b>0.97</b>
grc									1.00	0.28	-.11	0.45	0.31	0.31	0.31
irl										1.00	0.06	-.15	<b>1.00</b>	<b>1.00</b>	<b>1.00</b>
ita											1.00	-.68	0.04	0.04	0.04
lux												1.00	-.15	-.15	0.15
nld													1.00	<b>1.00</b>	<b>1.00</b>
prt														1.00	<b>1.00</b>
swe															1.00

MPL fluctuations between countries. In so doing, the moving-average-minimal-path-length (MAMPL) method leads us to a set of  $M = 4$  correlation matrices (one for each index), having the size  $N \times N$ , where  $N = 15$  is the number of countries under consideration here. For example, Table II, for FCE ( $\equiv s_2$ ). *Nota bene* the matrix is symmetric (half of the elements are shown) but not all elements are necessarily positive.

Let us suppose that we filter these (four) second correlation matrices in order to retain a few terms, those which lead us to build a network for which the weights (i.e. the correlation coefficients) are greater than e.g. 0.9. The correlation coefficients, e.g. in the case of the variable node  $s_2$  (FCE fluctuations), as given in Table II, are emphasized in bold for those  $\geq 0.9$  at each couple of function nodes. Due to the filtering, one can easily see that not all 15 countries have at least one “bold node”, i.e. are connected through the variable node  $s_2$  (FCE), but only nine of them, namely AUT, BEL, DEU, ESP, GBR, IRL, NLD, PRT, and SWE. In Fig. 1, these nine countries are connected through the variable node FCE (the dashed arrows).

The above procedure can be repeated for GDP, GCF, and NEX, whence obtaining the other “clusters” in Figs. 1 and 2, with respectively 9, 8 and 7 countries (for this filter value). Let us notice that GRC does not belong to any cluster\*.

The cluster contributions to the Hamiltonian can thus be the variable  $s_2$ . Then one obtains that the cost function  $\mathcal{H}$  associated to the factor graph (Figs. 1, 2) based on these four variables reads

\*GRC does not appear in the Hamiltonian, because (see Figs. 1, 2) GRC is *not* connected to the other countries through any  $s_i$  by a link having a weight greater than 0.9. If the linkage threshold is established to a lower value, e.g.  $|C| \geq 0.8$ , its function node appears as  $(\text{GRC})(s_1, s_4)$ , i.e. it belongs to the same clusters as Italy.

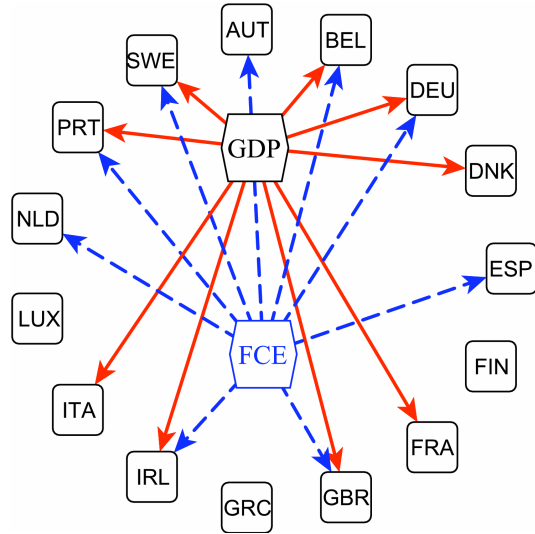


Fig. 1. The factor graph associated to the first 15 EU country connections, according to the strongest correlations in the GDP and FCE.

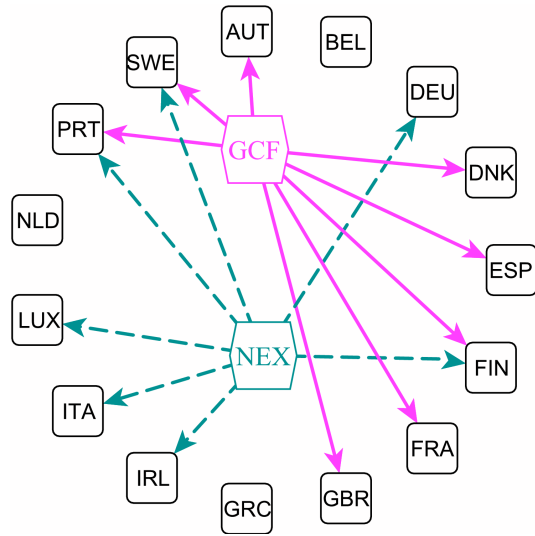


Fig. 2. The factor graph associated to the first 15 EU country connections, according to the strongest correlations in the GCF and NEX.

$$\mathcal{H} = \mathcal{H}_1 + \mathcal{H}_2 + \mathcal{H}_3 + \mathcal{H}_4,$$

where

$$\mathcal{H}_1 = (\text{LUX})(s_4) + (\text{NLD})(s_2),$$

$$\mathcal{H}_2 = (\text{ITA})(s_1, s_4) + (\text{AUT})(s_2, s_3) + (\text{BEL})(s_1, s_2) + (\text{DNK})(s_1, s_3),$$

$$\begin{aligned}\mathcal{H}_3 &= (\text{ESP})(s_2, s_3) + (\text{FIN})(s_3, s_4) + (\text{FRA})(s_1, s_3) + (\text{DEU})(s_1, s_2, s_4) \\ &\quad + (\text{GBR})(s_1, s_2, s_3) + (\text{IRL})(s_1, s_2, s_4), \\ \mathcal{H}_4 &= (\text{PRT})(s_1, s_2, s_3, s_4) + (\text{SWE})(s_1, s_2, s_3, s_4),\end{aligned}$$

from which one could write the equilibrium probability distribution, the partition function and the free energy, introduced here above.

#### 4. Discussion

Instead of writing a Hamiltonian as a function of the function nodes, let us project the dynamics of the factor graph into a phase space spanned by the variable nodes. Let us recall that a cluster  $\alpha$  was defined as a subset of the factor graph such that if a function node belongs to  $\alpha$ , then all the variable nodes  $s_\alpha$  also belong to  $\alpha$ . We can write all the possible combinations of the four variable nodes and find the Hamiltonian corresponding to function nodes. Let us take for example the combination ( $s_1 \equiv \text{GDP}$ ;  $s_2 \equiv \text{FCE}$ ;  $s_3 \equiv \text{GCF}$ ). Then, the function nodes connected *only* to these *three* variables (not necessarily to all of them) *and*

TABLE III

Clustering of the first 15 EU countries in a 4-variable factor graph approach after filtering (see text) and projecting in a 3-variable node phase space; the number of links in the cluster, the maximum possible number of links, subsequently the relevant ratio, and the entropy of each cluster are given.

Variable nodes	Cluster of function nodes	Number of links	Maximum number of links	Ratio	Entropy
GDP- -FCE- -GCF	-AUT-BEL- -DNK-ESP-FRA- -GBR-NLD-	14	28	0.500	0.347
GDP- -FCE- -NEX	-BEL-DEU- -IRL-ITA- -LUX-NLD	12	24	0.500	0.347
GDP- -GCF- -NEX	-DNK-FIN- -FRA- -ITA-LUX	9	20	0.450	0.359
FCE- -GCF- -NEX	-AUT-ESP- -FIN- -LUX-NLD	8	20	0.400	0.367

not to the fourth one ( $s_4 \equiv \text{NEX}$ ) are AUT, BEL, DNK, ESP, FRA, GBR, and NLD. This means a cluster that we can see in the first row in Table III. The same can be done for the other three combinations, leading to another set of clusters.

In so doing clustering [17] properties appear through e.g. an entropy, Eq. (7). The values are given in Table III for the clusters made of three variable nodes. As a not obvious consequence of this cluster analysis technique, it is observed that the *maximum* entropy (0.367) corresponds to the clustering which does *not* explicitly include the GDP but only the consumption, investment, and trade components. Another point can be deduced from the *minimal* entropy (0.347) clustering scheme, i.e. it is obtained from the coupling between GDP and FCE. The results confirm intuitive economic theory and practice expectations at least as regards geographical connexions. However, deep discussions of these findings are left for economists.

In conclusion, let us recall the frame of our work and our findings: relevant microscopic description of a system relies on a coarse-grained reduction of its internal variables. We have presented a way to do so for a bipartite graph having on one hand countries, on the other hand macroeconomy indicators. We have obtained a Hamiltonian description. The technique can of course be generalized and applied to many other socioeconomy networks.

## 5. Conclusion

Complex networks have become an active field of research in physics [18]. These systems are usually composed of a large number of internal components (the nodes and links), and describe a wide variety of systems of high intellectual and technological importance. Relevant questions pertain to the characterization of the networks. Investigations of the case of directed and/or weighted networks are not so common. The occurrence of community clustering for networks having nodes possessing a vector-like characteristics has been rarely studied. We have attempted to do so through a revival of some clustering variation method in the framework of some macroeconomy study.

We have taken as an example the weighted fully connected network of the  $N = 15$  first countries forming the European Union in 2005 (EU-25). The ties between countries are supposed to result (be proportional) to the degree of similitude of the macroeconomic fluctuations annual rates of *four* macroeconomic indicators, i.e. GDP, FCE, GCF, and NEX over *ca.* a 15 year time span.

Averaging the yearly increment correlations a weighted bipartite network has been built having the four “degrees of freedom” and the fifteen “countries” on the other hand as basis. The analysis shows the importance of  $N$ -body interactions in particular when observing the macroeconomy states as a function of time. This leads to identify and display clusters of countries, clusters resulting from projections onto a high-dimensional phase space spanned by indicators, taken as independent variables. This approach generalizes usual projection methods by accounting for the complex geometrical connections resulting from vector-like nodes.



In particular such a measure of collective habits does fit the usual and practical expectations defined by politicians, journalists, or economists, through so-called “common factors” [19, 20]. The analysis reveals geographical connexions indeed. It is expected that the technique can be applied to many types of physical and socioeconomic networks.

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